# The IRS Population Count: An Update 

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In a paper presented at the 1993 Annual Meetings of the American Statistical Association, the authors presented the results of their first attempt to use administrative records available at the Internal Revenue Service (IRS) to count the population of the United States (see Sailer, Weber, and Yau, 1993). In that paper, we noted that a major problem in this use of IRS administrative records was the presence in our files of information documents for deceased individuals. This was because several years could pass between the death of an individual and the closing out of all accounts listed in his or her name. In addition, we had some reason to be nervous about the accuracy of our gender coding, since it was based entirely on the interpretation of each individual's first name by some computer software we had developed. Poor reporting of social security numbers of dependents was a further obstacle to getting a correct count.

As will be discussed later, a number of these problems have been dealt with over the last five years, and it appeared to be an opportune time to research whether our processing changes had improved our ability to use IRS records for the purpose of counting the population. This paper covers the results of that research.

Organizationally, this paper is divided into four sections. First, we will demonstrate how administrative records can be used to compute a population estimate. Then we will discuss the reliability of this estimate. Next, we will compare estimates from our data base, classified by age, sex, and state, to population data published by the Census Bureau. And finally, we will summarize our conclusions and make some recommendations for further research.

## Computation of an IRS Administrative Records Population

Citizens and residents of the United States have numerous opportunities to come to the attention of the Internal Revenue Service. Obviously, the 61 percent of the population that files individual tax returns, either
as primary or secondary taxpayers, is easy enough to count. These individuals also report, as exemptions, any children or other individuals they are supporting.
In addition, individuals covered by salaries and wages are generally reported to the IRS on Forms W-2; individuals making contributions to Individual Retirement Arrangements (IRAs) or Simplified Employee Pension (SEP) accounts on Form 5498; individuals receiving gross distributions from IRAs, SEPs, or other pension plans on Forms 1099-R; recipients of interest on Forms 1099-INT; recipients of dividends on Forms 1099-DIV; recipients of original issue discounts on Forms 1099-OID; recipients of patronage dividends on Forms 1099-PATR; recipients of government transfer payments on Forms 1099-G; recipients of social security benefits on Forms SSA1099; sellers of capital assets on Forms 1099-B; sellers of real estate on Forms 1099-S; contractors with the Federal Government on Forms 8596; winners at gambling on Forms W-2G; payers of mortgage interest on Forms 1098; and recipients of many types of nonemployment compensation, including prizes, awards,

Table 1: Components of the IRS Population Count

| (Frequencies in 1,000's) | Weighted <br> Number | Cumulative |
| :--- | ---: | ---: |
| Primary Taxpayers (TY 1993) | 112,029 | 112,029 |
| Secondary taxpayers | 46,772 | 158,801 |
| Dependents without information docs | 45,868 | 204,669 |
| Non-filers with information docs | 45,257 | 249,926 |
| Dependents without SSNs | 6,674 | 256,600 |
| Deaths before July 1, 1994 | 4,331 | 252,269 |

rents, royalties, crop insurance payments, and golden parachute payments on Forms 1099-MISC.

Table 1 details how we used all of this information to count the population covered by IRS administrative records. We started, of course, with filers of tax returns for Tax Year 1993 (i.e., returns generally filed on or around April 15, 1994). However, contrary to our usual practice in our Statistics of Income reports, we did not count anybody filing a prior-year return in 1994, since these individuals had a chance of being captured as recipients of information documents. We also excluded anybody filing from a foreign address, since we wanted to compare our results to Census data for 1994, and Census does not count U.S. citizens living abroad. We counted 112.0 million current-year returns with U.S. addresses.

On joint returns selected for this sample, we counted the secondary taxpayers--a total of 46.8 million. This brought our count to 158.8 million.

We also counted dependents, but not all of them. Dependents with income could be picked up in our sample of information documents or in our sample of tax return filers, so initially we only counted those dependents who had SSNs, but for whom a search of our administrative records master files revealed no records. There were 45.9 million such dependents.

To the 204.7 million individuals counted thus far, we added 45.3 million non-filers with information documents. We got these individuals by pulling a simple, random sample of individuals with at least one information document on the Information Returns Master File, and then eliminating all who appeared either as a primary or a secondary taxpayer on a tax return. If they appeared on a tax return as a dependent, we left them in, since we were not including dependents with information documents in our count. Again, we eliminated any prior-year documents received by the IRS in 1994, and we did not count documents issued to individuals at foreign addresses.

Unfortunately, our file also contained 6.7 million dependents for whom no SSN was given. This was a major improvement over the 11.4 million dependents for whom no SSN was given for 1989, but still a disappoint-ment. Obviously, in the absence of an SSN, we could neither check the Information Returns Master File (IRMF) for income, nor the Year of Birth File for age. We did not have much choice but to count such dependents in the lowest age category, and assume that they were not information document recipients. From our Taxpayer Usage Study, [IRS, 1994-2], we know that an estimated 3.3 million taxpayers checked a box indicating that the dependent was under age 1 , and therefore not required to have an SSN. So, for nearly one-half of these dependents, we know we have the correct age. Luckily, this problem should pretty much disappear in future years, for IRS is no longer sending out refund checks to taxpayers who fail to provide dependent SSNs, or who provide non-verifiable ones.

At this point, our count is at 256.6 million. As mentioned previously, experience has taught us that some of the individuals in this count are deceased. Our big improvement this year was that the Social Security Administration was willing to share with IRS information they gather from various sources on which

SSNs belong to the deceased, including the owner's date of death. This meant that we no longer had to make case-by-case decisions as to who in our sample was alive on January 1, 1994--the date of the Census estimates we were using for comparison purposes. Anybody with a date of death prior to $1 / 1 / 94$ was simply taken out of the IRS count. This left with an IRS "population count" of 252.3 million, or 97.36 percent of the Census estimate of 259.1 million.

## Evaluation of the Estimate

The estimates presented in Table 1 are based on a highly stratified sample of 104,605 individual income tax returns [Internal Revenue Service, 1995], supplemented by a simple random sample of 45,257 individuals for whom our files contained information documents, but no tax returns. Therefore, the estimates are subject to sampling error. Our 95 percent confidence interval is between 251.0 and 253.5 million. So our estimate lies between 96.88 and 97.84 percent of the Census figure. At this point, it should also be noted that Census admits to an undercount of about 4 million individuals. Assuming that is correct, we have identified between 95.40 and 96.34 percent of the true population in our administrative records file.

The Census figures are updates of the counts from the 1990 Census, using data on births, deaths, immigration, and emigration [U.S. Bureau of the Census, 1998]. While they are not subject to sampling variability, they do contain non-sampling errors. The IRS data are subject to non-sampling error as well. While every effort has been made to eliminate incorrect SSNs and substitute correct ones where our historical files provided this information, it is quite likely that our sample still contains incorrect dependent SSNs. These could lead to false matches or false non-matches to information documents. Missing dependent SSNs could only lead to false non-matches, which would have the effect of overstating the IRS data, since we would have no way of detecting whether these dependents were already being counted in the "nonfilers with information documents" universe.

## Comparisons to Census

Let us now look at the age and sex distribution of individuals in our file of administrative records. As mentioned previously, age and sex were added to our file simply by matching to an extract from the Social Security Administration's (SSA) Year of Birth file, which IRS receives for administrative and research purposes. For those few individuals with missing or
invalid SSNs, the sex code was generated by matching the first name against a dictionary of gender-coded names. The age was imputed with the help of an algorithm that took into account the individual's sources of income (for example Social Security retirement income), the entry in the "over 65 " checkbox on the return, and, where available, the ages of the spouse and any dependent children shown on the same return.
Table 2. Number of individuals (in 1,000 's), Jan.1, 1994. IRS and Census estimates

| Total | IRS | Deaths by | Adj. IRS | Census | Adj. IRS as | Census | Adjusted | Adj. IRS as \% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age |  | 1/1/1994 |  |  | \% of Census | undercount | Census | of Adj. Census |
| Under 15 | 55,897 | 35 | 55,862 | 57,337 | 97.43 | 1,822 | 59,159 | 94.43 |
| 15 under 25 | 33,842 | 179 | 33,663 | 35,942 | 93.66 | 1,146 | 37,088 | 90.77 |
| 25 under 35 | 40,952 | 46 | 40,906 | 41,354 | 98.92 | 1,037 | 42,391 | 96.50 |
| 35 under 45 | 40,332 | 85 | 40,247 | 41,658 | 96.61 | 486 | 42,144 | 95.50 |
| 45 under 55 | 29,118 | 172 | 28,946 | 29,870 | 96.91 | 46 | 29,916 | 96.76 |
| 55 under 65 | 20,430 | 299 | 20,131 | 21,018 | 95.78 | -189 | 20,829 | 96.65 |
| 65 under 75 | 19,360 | 949 | 18,412 | 18,712 | 98.39 | -175 | 18,537 | 99.32 |
| 75 and over | 16,670 | 2,567 | 14,102 | 14,446 | 97.62 | -126 | 14,320 | 98.48 |
| Total | 256,600 | 4,331 | 252,269 | 260,337 | 96.90 | 4,047 | 264,384 | 95.42 |
|  |  |  |  |  |  |  |  |  |
| Male |  |  |  |  |  |  |  |  |
| Age |  |  |  |  |  |  |  |  |
| Under 15 | 28,448 | 16 | 28,433 | 29,353 | 96.86 | 927 | 30,280 | 93.90 |
| 15 under 25 | 17,396 | 115 | 17,281 | 18,347 | 94.19 | 603 | 18,950 | 91.19 |
| 25 under 35 | 21,181 | 36 | 21,145 | 20,677 | 102.26 | 618 | 21,295 | 99.29 |
| 35 under 45 | 20,311 | 65 | 20,246 | 20,648 | 98.05 | 325 | 20,973 | 96.53 |
| 45 under 55 | 14,505 | 120 | 14,386 | 14,591 | 98.59 | 59 | 14,650 | 98.19 |
| 55 under 65 | 9,955 | 193 | 9,762 | 9,984 | 97.78 | -63 | 9,921 | 98.40 |
| 65 under 75 | 8,714 | 595 | 8,119 | 8,290 | 97.94 | -51 | 8,239 | 98.54 |
| 75 and over | 6,294 | 1,184 | 5,110 | 5,185 | 98.55 | -21 | 5,164 | 98.95 |
| Total | 126,804 | 2,323 | 124,481 | 127,075 | 97.96 | 2,398 | 129,473 | 96.14 |
|  |  |  |  |  |  |  |  |  |
| Female |  |  |  |  |  |  |  |  |
| Age |  |  |  |  |  |  |  |  |
| Under 15 | 27,446 | 19 | 27,427 | 27,984 | 98.01 | 895 | 28,879 | 94.97 |
| 15 under 25 | 16,446 | 64 | 16,382 | 17,595 | 93.11 | 543 | 18,138 | 90.32 |
| 25 under 35 | 19,771 | 10 | 19,761 | 20,677 | 95.57 | 418 | 21,095 | 93.67 |
| 35 under 45 | 20,021 | 20 | 20,001 | 21,010 | 95.20 | 161 | 21,171 | 94.47 |
| 45 under 55 | 14,613 | 53 | 14,560 | 15,279 | 95.29 | -14 | 15,265 | 95.38 |
| 55 under 65 | 10,474 | 106 | 10,368 | 11,034 | 93.96 | -126 | 10,908 | 95.05 |
| 65 under 75 | 10,646 | 353 | 10,293 | 10,422 | 98.76 | -124 | 10,298 | 99.95 |
| 75 and over | 10,377 | 1,423 | 8,954 | 9,261 | 96.69 | -105 | 9,156 | 97.79 |
| Total | 129,794 | 2,048 | 127,746 | 133,262 | 95.86 | 1,648 | 134,910 | 94.69 |
|  |  |  |  |  |  |  |  |  |

As can be seen from Table 2, the overall correspondence between Census and administrative records data is extremely good-even better than it was for 1989. The over-estimation of the " 75 and over"
class has disappeared for 1993, thanks to the date-ofdeath information added to the data base. Some of the apparent difference between coverage of males and females (particularly in the age classes under 25) has been eliminated with the help of sex codes from SSA.

The only age/sex class in which IRS shows more individuals than are shown in the Census data is males age 25 to 35 . However, the IRS estimate is still slightly below the adjusted Census estimate, so IRS may have been able to account for some young males missed in the 1990 Census.

| State | Census adjusted | \% of Census Adjusted |  |
| :---: | :---: | :---: | :---: |
|  |  | Census | IRS |
| ALABAMA | 4,293 | 98.29 | 99.54 |
| ALASKA | 614 | 98.17 | 94.92 |
| ARIZONA | 4,169 | 97.85 | 92.66 |
| ARKANSAS | 2,495 | 98.32 | 92.58 |
| CALIFORNIA | 32,250 | 97.39 | 93.38 |
| COLORADO | 3,731 | 98.14 | 98.01 |
| CONNECTICUT | 3,296 | 99.35 | 97.74 |
| DELAWARE | 720 | 98.31 | 98.86 |
| DIST. OF COLUMBIA | 588 | 96.36 | 87.05 |
| FLORIDA | 14,218 | 98.17 | 96.38 |
| GEORGIA | 7,200 | 98.02 | 98.60 |
| HAWAII | 1,199 | 98.25 | 98.30 |
| IDAHO | 1,157 | 98.05 | 99.16 |
| ILLINOIS | 11,873 | 99.04 | 96.47 |
| INDIANA | 5,783 | 99.52 | 96.62 |
| IOW A | 2,843 | 99.60 | 94.64 |
| KANSAS | 2,568 | 99.32 | 99.67 |
| KENTUCKY | 3,889 | 98.44 | 92.38 |
| LOUISIANA | 4,410 | 97.87 | 93.67 |
| MAINE | 1,248 | 99.27 | 99.10 |
| MARYLAND | 5,101 | 98.02 | 94.64 |
| MASSACHUSETTS | 6,070 | 99.52 | 95.51 |
| MICHIGAN | 9,558 | 99.31 | 92.25 |
| MINNESOTA | 4,587 | 99.58 | 97.17 |
| MISSISSIPPI | 2,726 | 97.93 | 94.59 |
| MISSOURI | 5,311 | 99.40 | 96.17 |
| MONTANA | 875 | 97.80 | 88.79 |
| NEBRASKA | 1,634 | 99.37 | 96.91 |
| NEVADA | 1,491 | 98.06 | 97.11 |
| NEW HAMPSHIRE | 1,144 | 99.19 | 99.28 |
| NEW JERSEY | 7,948 | 99.43 | 100.10 |
| NEW MEXICO | 1,704 | 97.15 | 103.54 |
| NEW YORK | 18,432 | 98.49 | 91.56 |
| NORTH CAROLINA | 7,196 | 98.25 | 97.14 |
| NORTH DAKOTA | 643 | 99.34 | 94.83 |
| OHIO | 11,179 | 99.33 | 95.73 |
| OKLAHOMA | 3,314 | 98.27 | 91.29 |
| OREGON | 3,141 | 98.28 | 98.33 |
| PENNSYLVANIA | 12,098 | 99.71 | 97.17 |
| RHODE ISLAND | 996 | 99.85 | 94.36 |
| SOUTH CAROLINA | 3,716 | 98.04 | 92.98 |
| SOUTH DAKOTA | 730 | 99.06 | 96.69 |
| TENNESSEE | 5,263 | 98.34 | 93.45 |
| TEXAS | 18,899 | 97.43 | 94.81 |
| UTAH | 1,939 | 98.44 | 94.92 |
| VERMONT | 586 | 98.92 | 102.68 |
| VIRGINIA | 6,677 | 98.11 | 96.62 |
| WASHINGTON | 5,430 | 98.31 | 98.26 |
| WEST VIRGINIA | 1,850 | 98.60 | 93.44 |
| WISCONSIN | 5,113 | 99.41 | 97.84 |
| W YOMING | 486 | 97.93 | 91.64 |
| Total | 264,384 | 98.47 | 95.42 |

If the IRS administrative data are to be used in a meaningful way to help Census identify individuals missing from the decennial Census, or even just to make intercensal estimates, it is important that they be
classifiable by geographic code. IRS data are somewhat problematical in this regard. Some IRS tax return addresses do not locate the taxpayer's residence-they may represent a tax accountant's address, a business address, a post office box in another town, or a rural route that crosses county lines. The addition of information documents to the data base provides the user with alternative addresses for each taxpayer-unfortunately, in some cases, with several alternatives. With a good deal of research, it may be possible to rank various types of information documents as to their likelihood of showing a residential address. It may also be possible to write algorithms that detect and eliminate addresses which are not residential.

Having admitted these shortcomings, we hasten to add that the vast majority of tax documents do contain addresses which can be used to code the residence of taxpayers. Unfortunately, the data base with which we were working was not designed to produce accurate estimates below the national level. Even at the state level, the estimates tend to show a good deal of sampling error. In order to minimize the error, we derived state estimates through a three-step process: number of primary taxpayers was taken straight from an IRS Master File Tabulation [Internal Revenue Service, 1994-1] (i.e., it is not subject to sampling error). The count of secondary taxpayers and dependents without income was ratio adjusted by the same percentage as number of primary taxpayers (i.e., it is subject to non-sampling error); and the non-filer population was left unadjusted (i.e., it is subject to sampling error). Table 3 shows the comparison of these estimates to Census population figures (adjusted for the undercount) by state. Also shown is a comparison of the official Census count to the adjusted Census count. It shows that, for nine states, the IRS estimate is actually closer to the adjusted Census population than is the official Census figure. For 16 more, the Census and IRS estimates are within three percentage points of one-another. The low coverage by IRS for New York ( 91.56 percent) and California (93.38) should not be a sampling variability problem, but may be related to their high rates of immigration. It may take immigrants a while to get into the IRS document systems.

## Conclusions and Recommendations

In commenting on the authors' earlier paper on this subject, John Czajka et al. wrote that we had "demonstrated the IRS administrative record system provides sufficiently high coverage of the U.S. resident population to be credible as the principal source of data
for an enumeration ... with administrative records" [Czajka et al., 1997]. We feel that the 1993 data, with improved SSN reporting, better gender codes, and addition of date of death information, reconfirm that conclusion. At very least, the Census Bureau, which has access to both tax return and information document files, should be studying how to use this information to identify individuals at addresses missed in the regular decennial Census. Much of the additional research that needs to be done-especially on the quality of address information on various types of documentswill have to be done at the Census Bureau, since only they can start with matched files of Census and administrative data. For our part, we will monitor the effects of improved SSN reporting on this type of data analysis.

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